

Normal distribution  
Exponential distribution  
Gamma distribution

**probability I**

**1398/ 2**

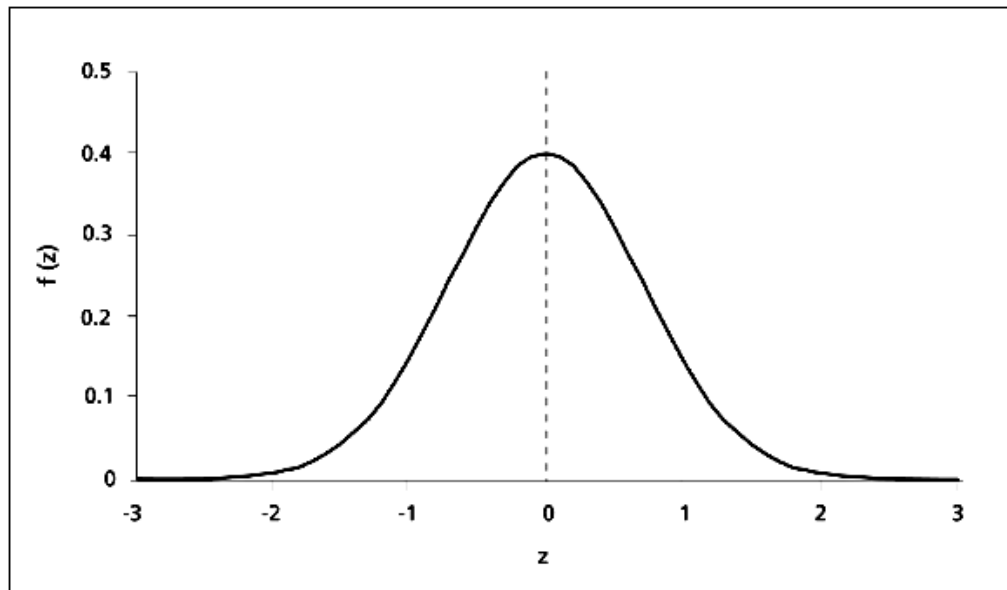
**D. Shahsavani**

# Standard Normal distribution

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A continuous random variable  $Z$  has **standard normal distribution** if

$$f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2} \quad z \in \mathbb{R}$$



notation

$$Z \sim N(0,1)$$

$$M(t) = E(e^{tX}) = e^{t^2/2}$$

$$E(Z) = \int_{-\infty}^{\infty} z \frac{1}{\sqrt{2\pi}} e^{-z^2/2} dz = 0$$

$$V(Z) = E(Z^2) = 1$$

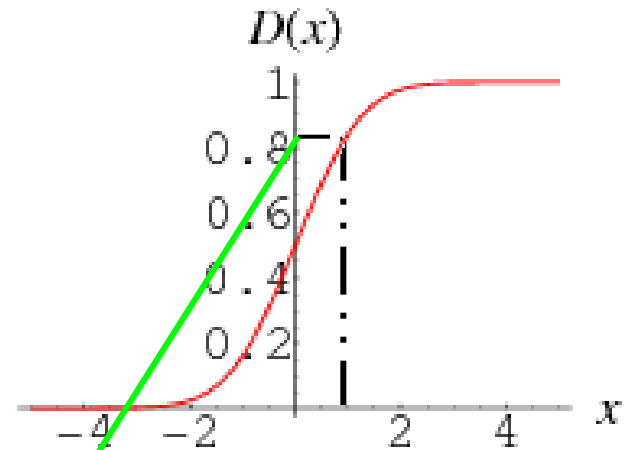
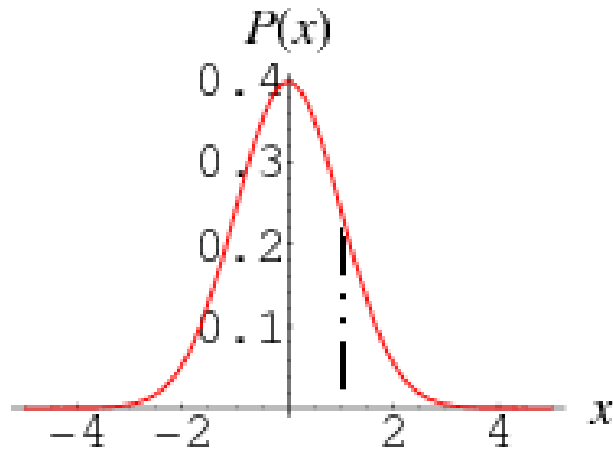
# Standard Normal distribution

## Cumulative Distribution function

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$$f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2} \quad z \in \mathbb{R}$$

$$\varphi(\alpha) = P(Z \leq \alpha) = \int_{-\infty}^{\alpha} \frac{1}{\sqrt{2\pi}} e^{-z^2/2} dz$$



$$\varphi(1) = P(Z \leq 1) = \int_{-\infty}^1 \frac{1}{\sqrt{2\pi}} e^{-z^2/2} dz$$

# Standard Normal Cumulative Probability Table



Cumulative probabilities for POSITIVE z-values are shown in the following table:

| z   | 0.00   | 0.01   | 0.02   | 0.03   | 0.04   | 0.05   | 0.06   | 0.07   | 0.08   | 0.09   |
|-----|--------|--------|--------|--------|--------|--------|--------|--------|--------|--------|
| 0.0 | 0.5000 | 0.5040 | 0.5080 | 0.5120 | 0.5160 | 0.5199 | 0.5239 | 0.5279 | 0.5319 | 0.5359 |
| 0.1 | 0.5398 | 0.5438 | 0.5478 | 0.5517 | 0.5557 | 0.5596 | 0.5636 | 0.5675 | 0.5714 | 0.5753 |
| 0.2 | 0.5793 | 0.5832 | 0.5871 | 0.5910 | 0.5948 | 0.5987 | 0.6026 | 0.6064 | 0.6103 | 0.6141 |
| 0.3 | 0.6179 | 0.6217 | 0.6255 | 0.6293 | 0.6331 | 0.6368 | 0.6406 | 0.6443 | 0.6480 | 0.6517 |
| 0.4 | 0.6554 | 0.6591 | 0.6628 | 0.6664 | 0.6700 | 0.6736 | 0.6772 | 0.6808 | 0.6844 | 0.6879 |
| 0.5 | 0.6915 | 0.6950 | 0.6985 | 0.7019 | 0.7054 | 0.7088 | 0.7123 | 0.7157 | 0.7190 | 0.7224 |
| 0.6 | 0.7257 | 0.7291 | 0.7324 | 0.7357 | 0.7389 | 0.7422 | 0.7454 | 0.7486 | 0.7517 | 0.7549 |
| 0.7 | 0.7580 | 0.7611 | 0.7642 | 0.7673 | 0.7704 | 0.7734 | 0.7764 | 0.7794 | 0.7823 | 0.7852 |
| 0.8 | 0.7881 | 0.7910 | 0.7939 | 0.7967 | 0.7995 | 0.8023 | 0.8051 | 0.8078 | 0.8106 | 0.8133 |
| 0.9 | 0.8159 | 0.8186 | 0.8212 | 0.8238 | 0.8264 | 0.8289 | 0.8315 | 0.8340 | 0.8365 | 0.8389 |
| 1.0 | 0.8413 | 0.8438 | 0.8461 | 0.8485 | 0.8508 | 0.8531 | 0.8554 | 0.8577 | 0.8599 | 0.8621 |
| 1.1 | 0.8643 | 0.8665 | 0.8686 | 0.8708 | 0.8729 | 0.8749 | 0.8770 | 0.8790 | 0.8810 | 0.8830 |
| 1.2 | 0.8849 | 0.8869 | 0.8888 | 0.8907 | 0.8925 | 0.8944 | 0.8962 | 0.8980 | 0.8997 | 0.9015 |
| 1.3 | 0.9032 | 0.9049 | 0.9066 | 0.9082 | 0.9099 | 0.9115 | 0.9131 | 0.9147 | 0.9162 | 0.9177 |
| 1.4 | 0.9192 | 0.9207 | 0.9222 | 0.9236 | 0.9251 | 0.9265 | 0.9279 | 0.9292 | 0.9306 | 0.9319 |
| 1.5 | 0.9332 | 0.9345 | 0.9357 | 0.9370 | 0.9382 | 0.9394 | 0.9406 | 0.9418 | 0.9429 | 0.9441 |
| 1.6 | 0.9452 | 0.9463 | 0.9474 | 0.9484 | 0.9495 | 0.9505 | 0.9515 | 0.9525 | 0.9535 | 0.9545 |
| 1.7 | 0.9554 | 0.9564 | 0.9573 | 0.9582 | 0.9591 | 0.9599 | 0.9608 | 0.9616 | 0.9625 | 0.9633 |
| 1.8 | 0.9641 | 0.9649 | 0.9656 | 0.9664 | 0.9671 | 0.9678 | 0.9686 | 0.9693 | 0.9699 | 0.9706 |
| 1.9 | 0.9713 | 0.9719 | 0.9726 | 0.9732 | 0.9738 | 0.9744 | 0.9750 | 0.9756 | 0.9761 | 0.9767 |
| 2.0 | 0.9772 | 0.9778 | 0.9783 | 0.9788 | 0.9793 | 0.9798 | 0.9803 | 0.9808 | 0.9812 | 0.9817 |

$$\varphi(0) = P(Z \leq 0) = 0.5 \quad \varphi(1.32) = P(Z \leq 1.32) = 0.9066$$

$$\varphi(-\alpha) = 1 - \varphi(\alpha) \quad \varphi(-1.32) = 1 - \varphi(1.32) = 1 - 0.9066 = 0.0934$$



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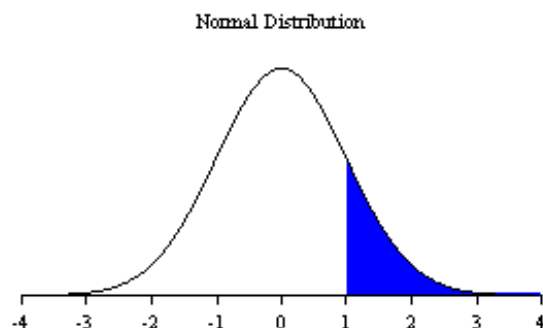
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## Normal Distribution

This page contains calculators for the normal distribution. The first calculator allows to enter the mean, standard deviation, and cutoff points; the area under the normal distribution is computed. The second calculator does just the reverse. These calculators can be used instead of a table of the normal distribution.

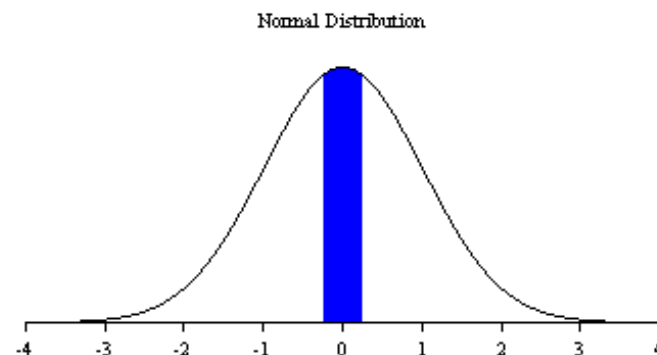
Choose parameters and the area will be computed.



Mean:  Sd:

- Above   
 Below   
 Between  and   
 Outside  or

Shaded area: 0.158655



Mean:  Sd:

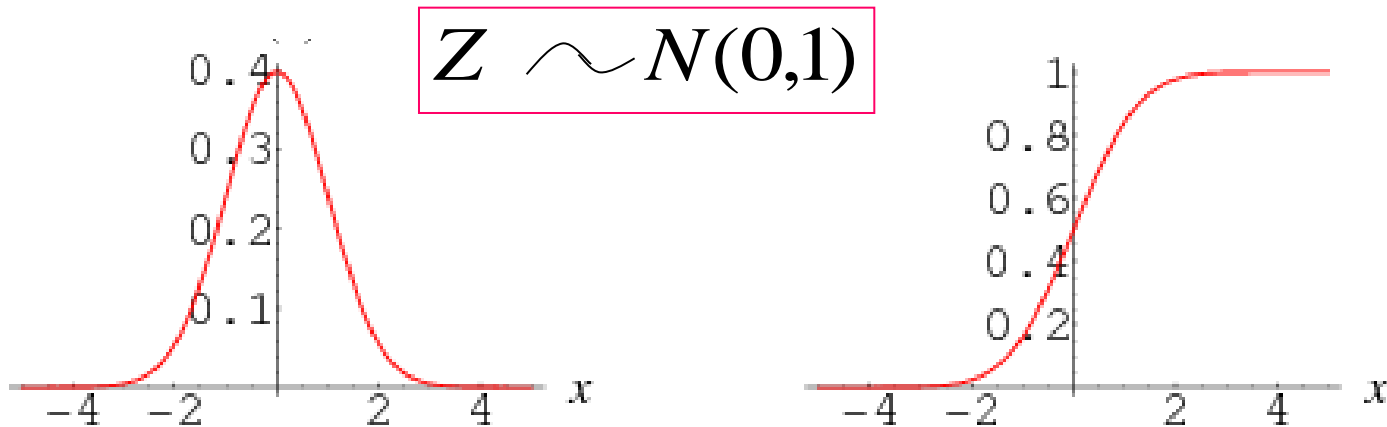
Shaded Area:

- Above  
 Below  
 Between:  and   
 Outside

# Standard Normal distribution

Reminder:

$$f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2} \quad z \in \mathbb{R} \quad \varphi(\alpha) = P(Z \leq \alpha) = \int_{-\infty}^{\alpha} \frac{1}{\sqrt{2\pi}} e^{-z^2/2} dz$$



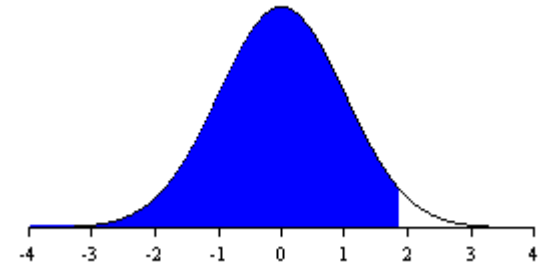
$$E(Z) = \int_{-\infty}^{\infty} z \frac{1}{\sqrt{2\pi}} e^{-z^2/2} dz = 0$$

$$V(Z) = E(Z^2) = 1$$

$$M(t) = E(e^{tX}) = e^{t^2/2}$$

# Normal distribution and Excel

$$P(Z \leq 1.87) = \varphi(1.87) = \int_{-\infty}^{1.87} \frac{1}{\sqrt{2\pi}} e^{-z^2/2} dz = 0.96925$$



Calculation of standard normal probabilities using **Excel**

**Function Arguments**

**NORMDIST**

**X** 1.87 = 1.87

**Mean** 0 = 0

**Standard\_dev** 1  $Z \sim N(0,1)$  = 1

**Cumulative** 1 = TRUE

= 0.969258091

Returns the normal cumulative distribution for the specified mean and standard deviation.

**Cumulative** is a logical value: for the cumulative distribution function, use TRUE; for the probability mass function, use FALSE.

Formula result = 0.969258091

[Help on this function](#)

OK Cancel

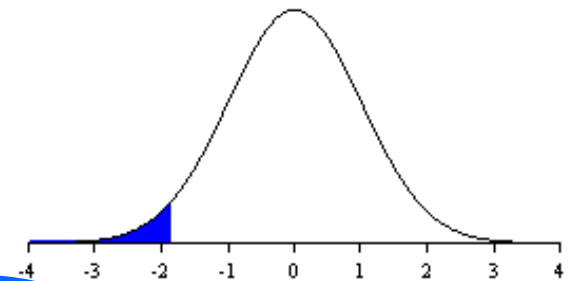
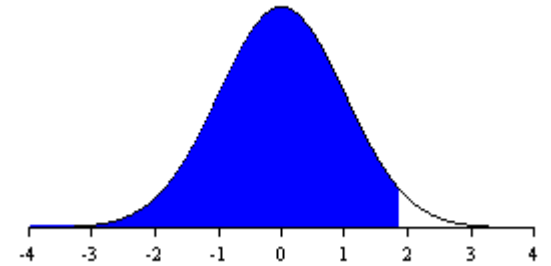
# calculation of probabilities

- $P(Z \leq 1.87) = \varphi(1.87) = \int_{-\infty}^{1.87} \frac{1}{\sqrt{2\pi}} e^{-z^2/2} dz = 0.96925$

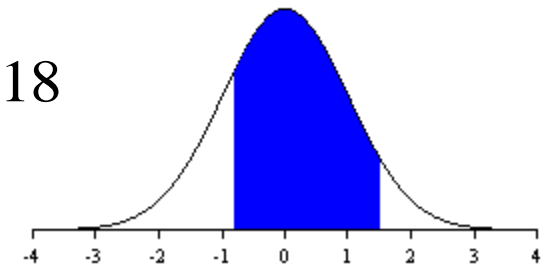
- $\varphi(-1.87) = P(Z \leq -1.87) = \int_{-\infty}^{-1.87} \frac{1}{\sqrt{2\pi}} e^{-z^2/2} dz = 0.0307$

$$= \int_{1.87}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-z^2/2} dz$$

$$= 1 - \int_{-\infty}^{1.87} \frac{1}{\sqrt{2\pi}} e^{-z^2/2} dz = 1 - \varphi(1.87)$$



Normal Distribution



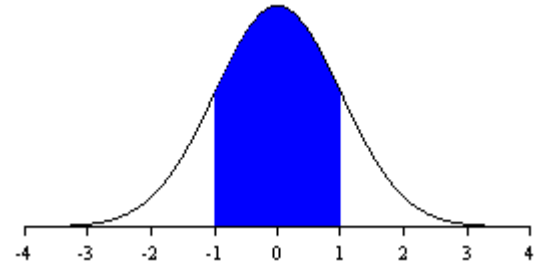
- $P(-0.8 < Z \leq 1.5) = \varphi(1.5) - \varphi(-0.8) = 0.9331 - 0.2118 = 0.7213$



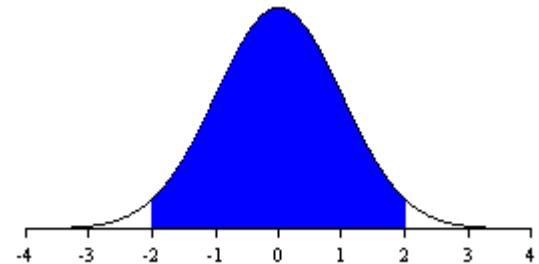
# Normal distribution

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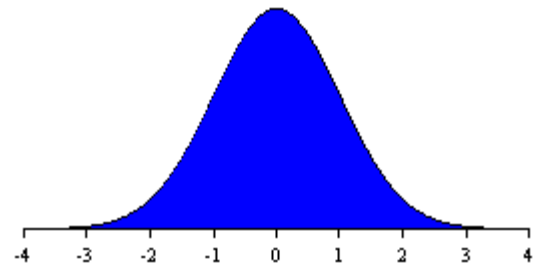
$$P(-1 < Z \leq 1) = \varphi(1) - \varphi(-1) = 0.68$$



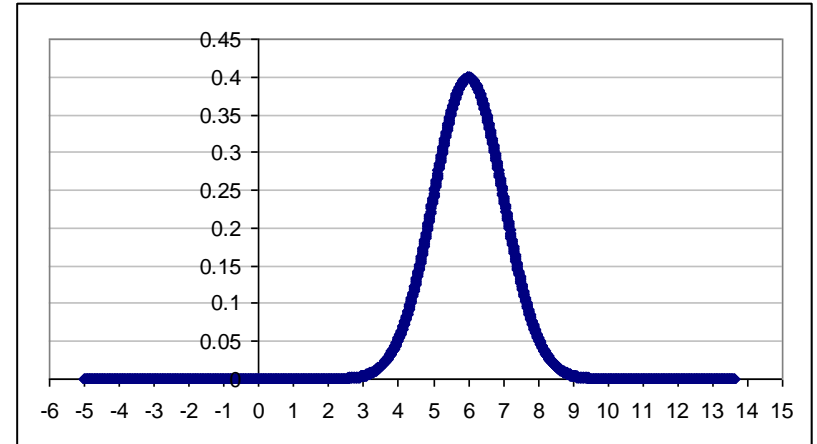
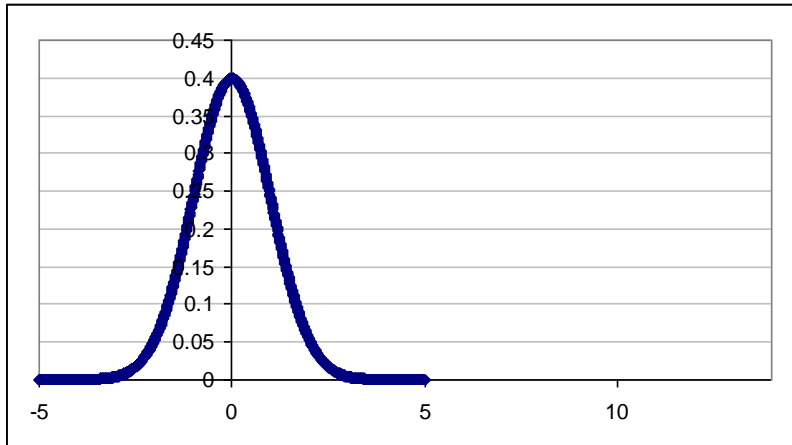
$$P(-2 < Z \leq 2) = \varphi(2) - \varphi(-2) = 0.95$$



$$P(-3 < Z \leq 3) = \varphi(3) - \varphi(-3) = 0.99$$



# Normal dist.- different means



$$f(z) = (\sqrt{2\pi})^{-1} e^{-\frac{1}{2}z^2}$$

**Standard Normal**

$$E(Z)=0$$

$$V(X)=1$$

$$Z \sim N(0,1)$$

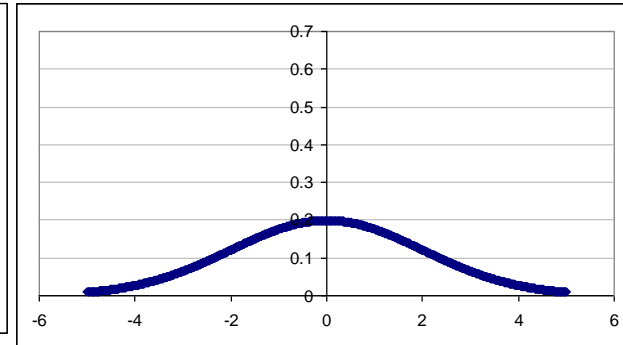
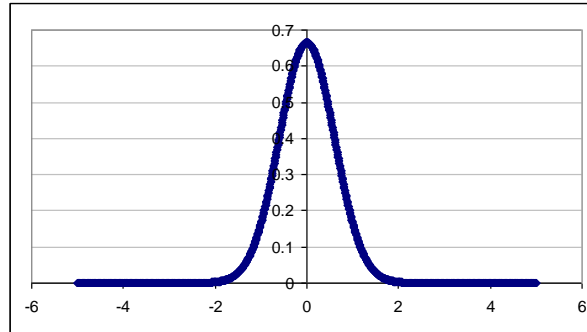
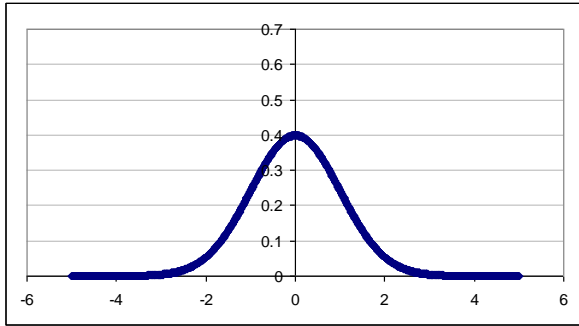
$$f(x) = (\sqrt{2\pi})^{-1} e^{-\frac{1}{2}(x-6)^2}$$

$$E(X)=6$$

$$V(X)=1$$

$$Z \sim N(6,1)$$

# Normal dist. – different variances



$$f(z) = (\sqrt{2\pi})^{-1} e^{-\frac{1}{2}z^2}$$

Standard Normal

$$E(Z)=0$$

$$V(X)=1$$

$$Z \sim N(0,1)$$

$$(0.6\sqrt{2\pi})^{-1} e^{-\frac{1}{2}\left(\frac{x}{0.6}\right)^2}$$

$$E(X)=0$$

$$V(X)=(0.6)^2$$

$$X \sim N(0,0.6^2)$$

$$(2\sqrt{2\pi})^{-1} e^{-\frac{1}{2}\left(\frac{y}{2}\right)^2}$$

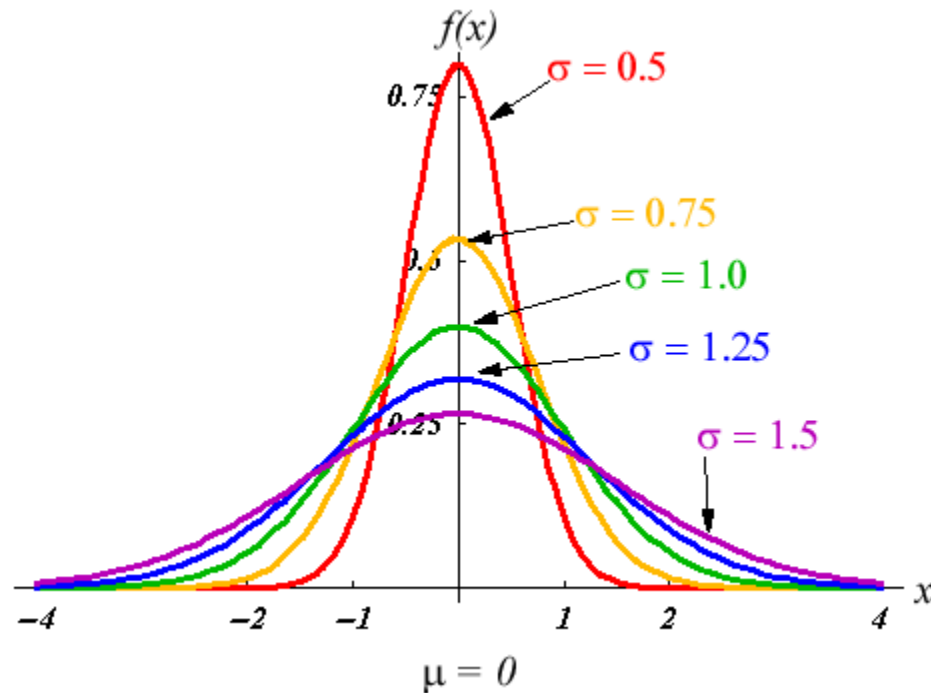
$$E(X)=0$$

$$V(X)=(2)^2$$

$$Y \sim N(0,2^2)$$

# Normal distribution

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# Normal distribution

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A continuous random variable  $X$  is called **normal random variable** if it has a probability density function of the form

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} \quad x \in \mathbb{R} \quad \mu \in \mathbb{R} \quad \sigma > 0$$

where  $\mu$  and  $\sigma^2$  are **mean** and **variance** of  $X$ , respectively. notation  $X \sim N(\mu, \sigma^2)$

Prove that:  $M(t) = e^{t\mu + t^2\sigma^2/2}$

$$M'(0) = \mu = E(X)$$

$$M''(0) = \mu^2 + \sigma^2$$

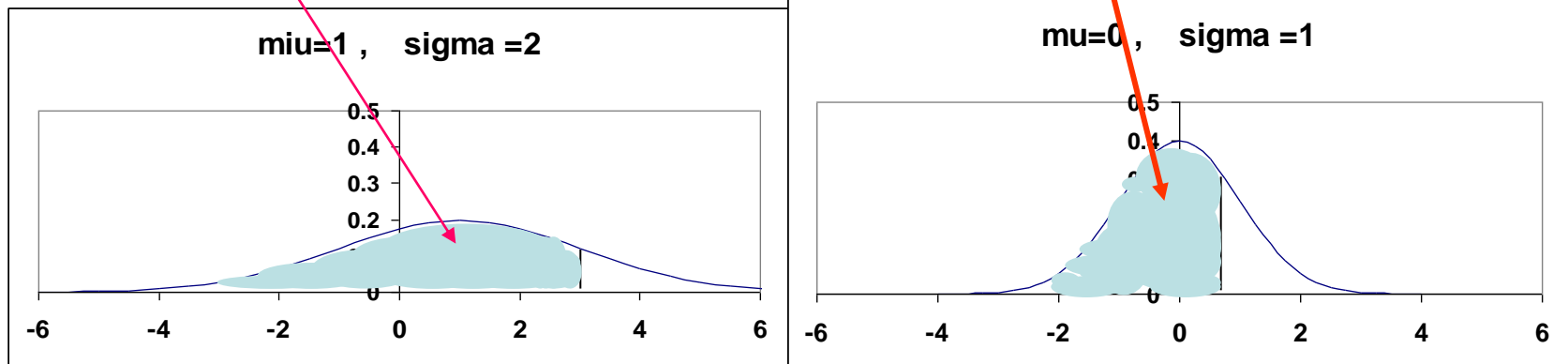
$$\int_{-\infty}^{+\infty} f(x) dx = 1$$

# Relationship

if  $X \sim N(\mu, \sigma^2)$  then  $[(X - \mu) / \sigma] \sim N(0,1)$

Suppose that:  $X \sim N(1,4)$

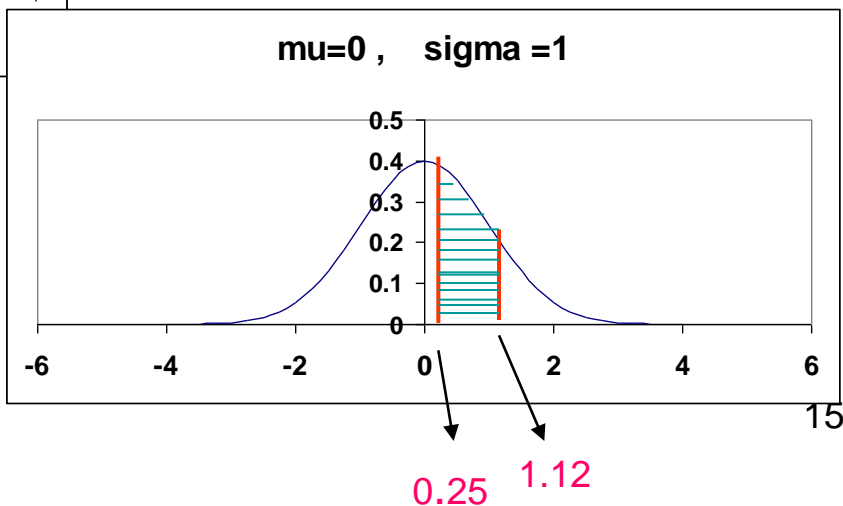
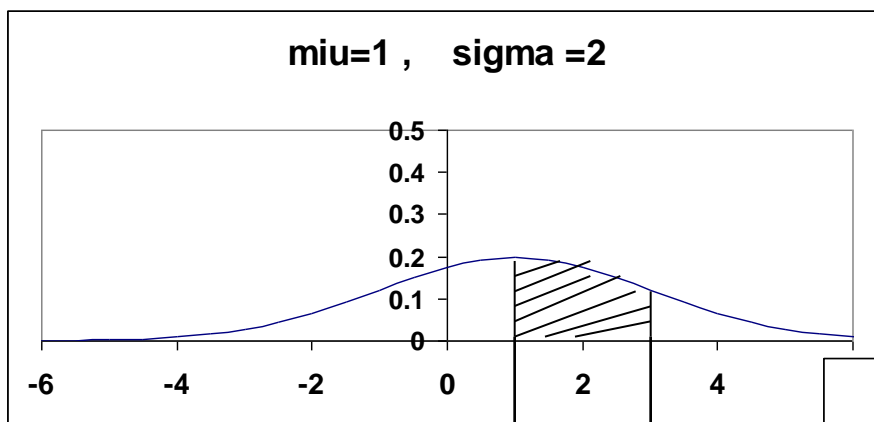
$$P(X \leq 3.24) = P\left(\frac{X - 1}{2} \leq \frac{3.24 - 1}{2}\right) = P(Z \leq 1.12) = \Phi(1.12) = 0.8686$$



# Example 2

Let  $X$  be normally distributed with parameters  $\mu=1$  and  $\sigma^2=4$ . Then

$$\begin{aligned} P(1.5 \leq X \leq 3.24) &= P\left(\frac{1.5-1}{2} \leq \frac{X-1}{2} \leq \frac{3.24-1}{2}\right) = P(0.25 \leq Z \leq 1.12) \\ &= \varphi(1.12) - \varphi(0.25) = 0.8686 - 0.5987 \end{aligned}$$

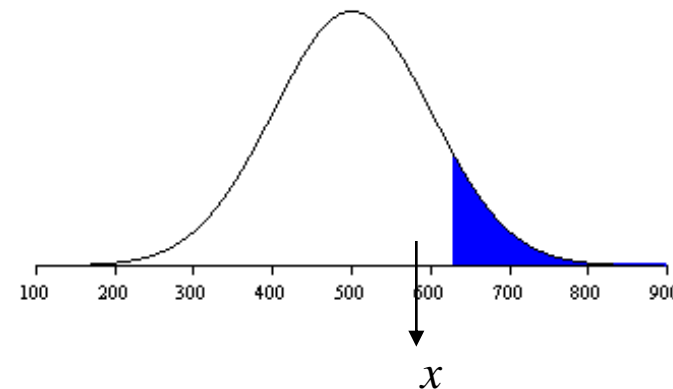


# Normal probability - Example 2

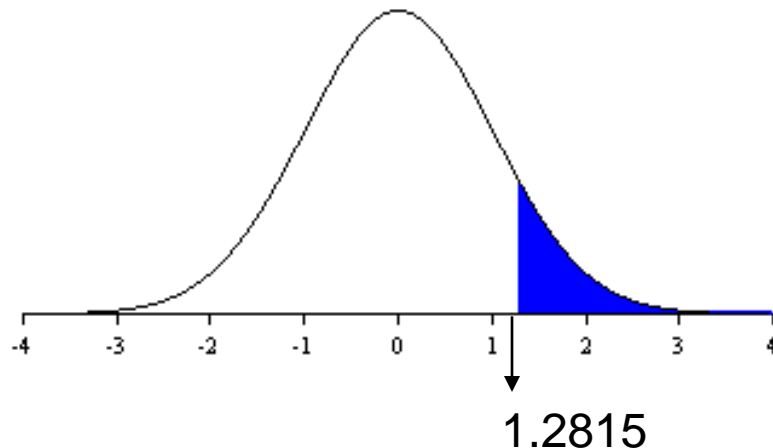
The score of a test are normally distributed with mean=500 and standard deviation=100. What is the score of a student that can be regarded as 10% of excellent students.  $X$ : scores of students in the exam :  $N(500, 100^2)$

We need an  $x$  such that  $P(X > x) = 0.1$

$$P(X \leq x) = 0.9 \Rightarrow P\left(\frac{X - 500}{100} \leq \frac{x - 500}{100}\right) = 0.9$$
$$P(Z \leq \frac{x - 500}{100}) = 0.9$$



$$P(Z \leq 1.2815) = \phi(1.2815) = 0.9$$



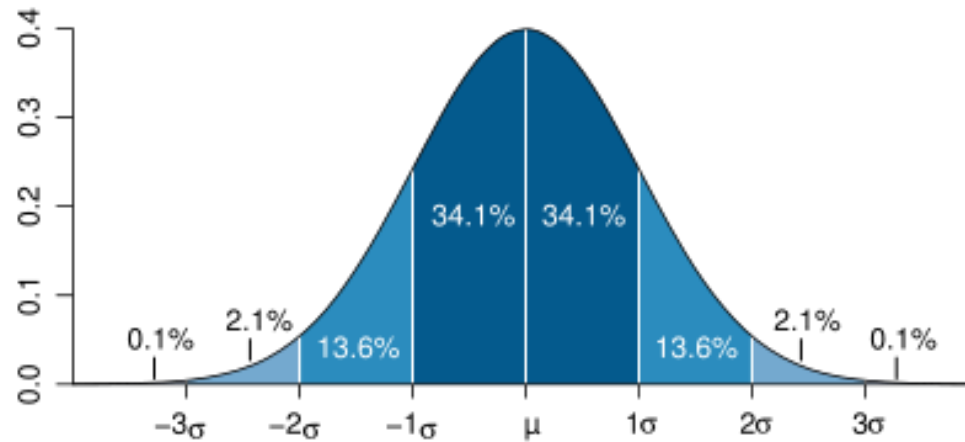
$$\Rightarrow \frac{x - 500}{100} \approx 1.28 \Rightarrow x \approx 628$$

A student should gain 628 scores that can be regarded as upper 10% of students



# Percent of value within intervals

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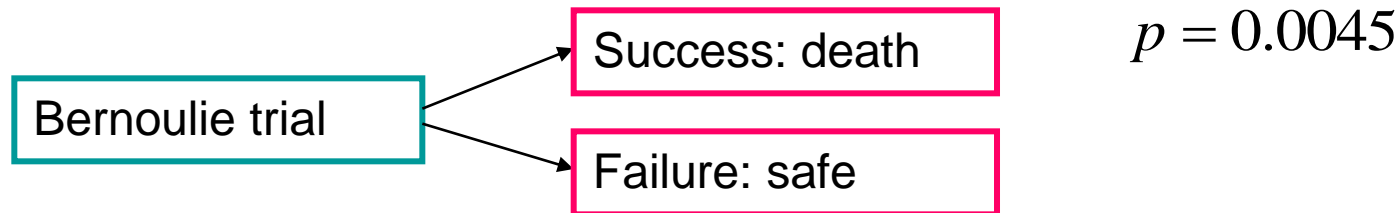
$$\begin{aligned} P(\mu \leq X \leq \mu + \sigma) &= P(\mu - \mu \leq X - \mu \leq \mu + \sigma - \mu) = P\left(\frac{0}{\sigma} \leq \frac{X - \mu}{\sigma} \leq \frac{\sigma}{\sigma}\right) \\ &= P(0 \leq Z \leq 1) = \varphi(1) - \varphi(0) = 0.8413 - 0.5 = 0.3413 \end{aligned}$$

$$\begin{aligned} P(\mu + \sigma \leq X \leq \mu + 2\sigma) &= P\left(1 \leq \frac{X - \mu}{\sigma} \leq 2\right) \\ &= P(1 \leq Z \leq 2) = \varphi(2) - \varphi(1) = 0.9772 - 0.8413 = 0.1353 \end{aligned}$$

# Discrete approximation – Poisson

Example : The infant mortality is 4.5 per 1000 live birth. Select 500 ( randomly) live birth, what is the probability of there are

- a) No infant death      b) at most three infant death



$X$  = Number of infant death in 500 live birth

$$X \sim B(500, 0.0045) ;$$

$$X \approx P(np) ; \quad P(X = x) \approx \frac{e^{-np} (np)^x}{x!}$$

$$P(X = x) = \binom{500}{x} (0.0045)^x (1 - 0.0045)^{500-x}$$

$$a) \quad P(X = 0) = (1 - 0.0045)^{500} \approx e^{-2.25} (2.25)^0 / 0! = 0.105$$

$$b) \quad P(X \leq 3) = \sum_{x=0}^3 \binom{500}{x} (0.0045)^x (1 - 0.0045)^{500-x} \approx \sum_{x=0}^3 e^{-2.25} (2.25)^x / x! = 0.809$$

# Continuous approximation

## Normal distribution

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Theorem: if  $X \sim B(n, p)$  ; then

$$\lim_{n \rightarrow \infty} P\left(a < \frac{X - np}{\sqrt{npq}} < b\right) = \frac{1}{\sqrt{2\pi}} \int_a^b e^{-x^2/2} dx$$

In the previous example

$$P(X \leq 3) = P(0 \leq X \leq 3) = P\left(\frac{0 - 2.25}{\sqrt{2,23}} \leq \frac{X - 2.25}{\sqrt{2,23}} \leq \frac{3 - 2.25}{\sqrt{2,23}}\right)$$

$$P\left(-1.5 \leq \frac{X - 2.25}{\sqrt{2,23}} \leq 0.5\right) \approx \frac{1}{\sqrt{2\pi}} \int_{-1.5}^{0.5} e^{-x^2/2} dx = ?$$

**But** this integral is only calculated numerically. It can be proved that (analytically)

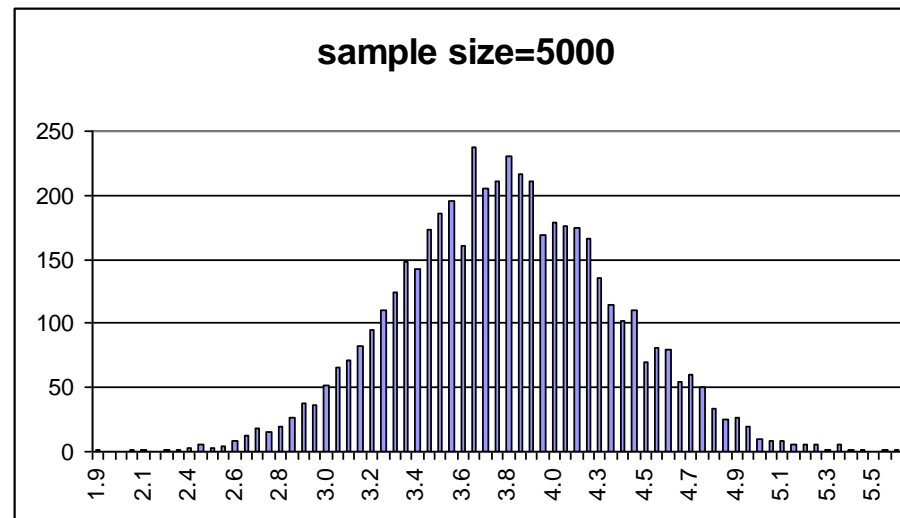
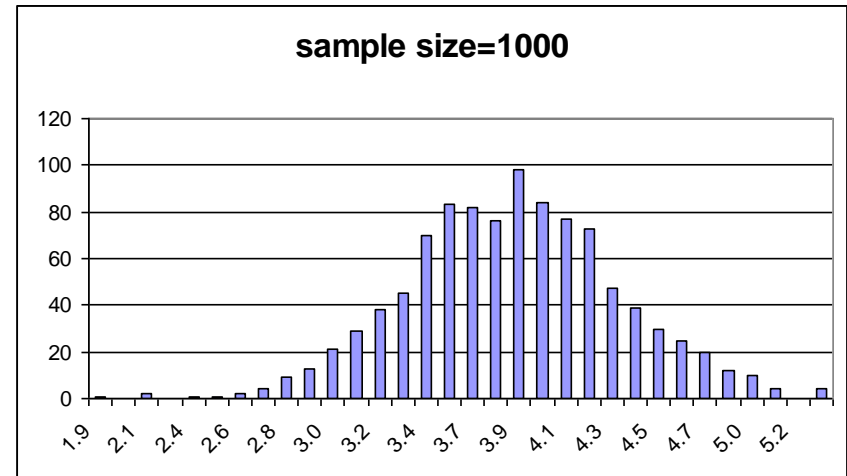
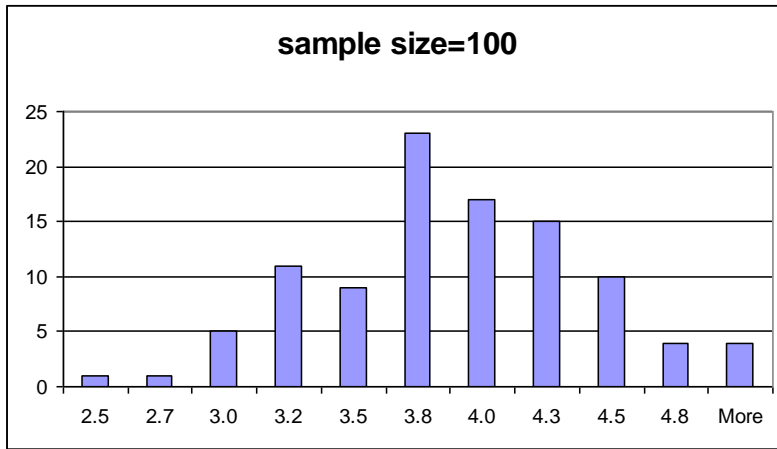
$$\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-x^2/2} dx = 1$$

**Note:**  $f(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$

can be regarded as a **probability density function** 19

# Distribution of infant's heights

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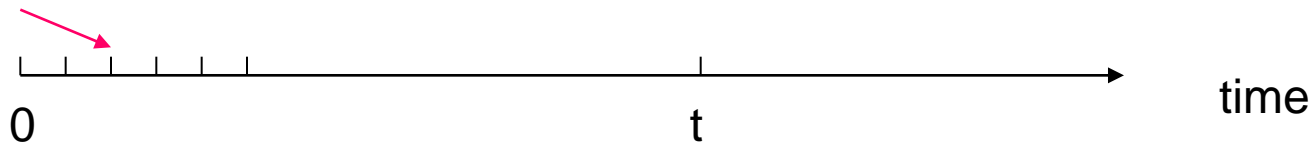


# Exponential distribution

Consider the events that are happened in a period of time between 0 and t. We saw that the number of these events ( let say N) can be modelled by **Poisson** distribution with parameter **lambda**, that is the average number of events in that period

$$P(N = n) = \frac{e^{-\lambda} \lambda^n}{n!} \quad n = 0, 1, 2, \dots$$

Unit of time



If **m**= the average number of events in each unit of time then

$$P(N = n) = \frac{e^{-mt} (mt)^n}{n!} \quad n = 0, 1, 2, \dots$$

Define X= **happening time of the first event**

$$P(X > t) = P(N = 0) = e^{-mt} \quad \text{or} \quad P(X < t) = 1 - e^{-mt}$$

$$F_X(t) = P(X \leq t) = \begin{cases} 0 & t < 0 \\ 1 - e^{-mt} & t > 0 \end{cases}$$

$$f(t) = \begin{cases} 0 & t < 0 \\ me^{-mt} & t \geq 0 \end{cases}$$

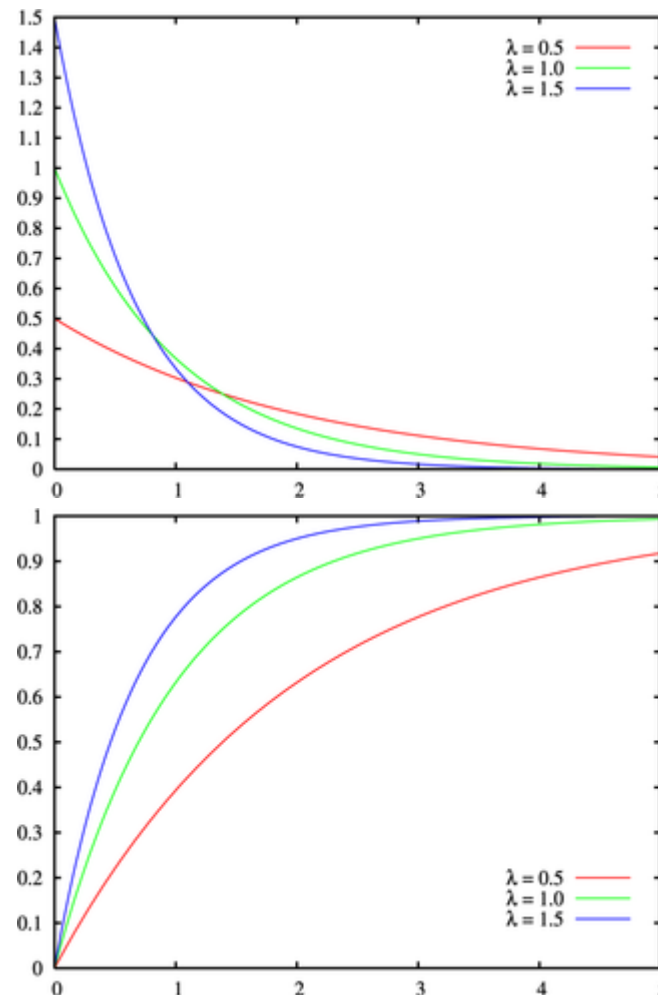
# Exponential distribution

A continuous random variable  $X$  has **Exponential distribution** with parameter lambda if

$$f(x) = \begin{cases} \frac{1}{\lambda} e^{-\frac{x}{\lambda}} & x > 0 \\ 0 & o.w. \end{cases}$$

$$F(x) = \begin{cases} 0 & x < 0 \\ 1 - e^{-\frac{x}{\lambda}} & x \geq 0 \end{cases}$$

Find  $E(x)$ ,  $\text{var}(X)$ ,  $M(t)$

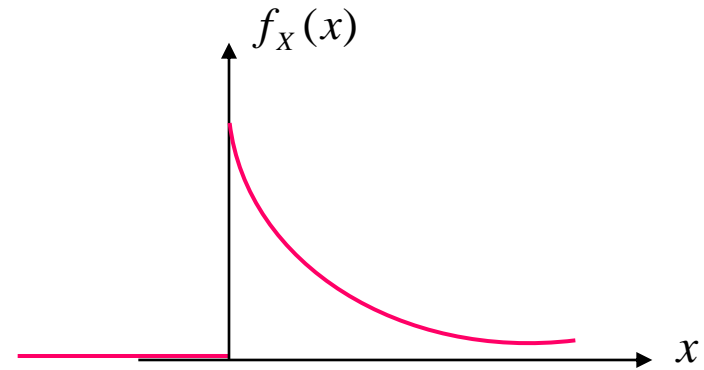


# Example

**X: the life time of an electrical component**

$$f_X(x) = \begin{cases} 0.01e^{-\frac{x}{100}} & x > 0 \\ 0 & \text{o.w.} \end{cases}$$

$$F(a) = P(X \leq a) = \int_{-\infty}^a f_X(t) dt = ?$$



$$F(t) = \begin{cases} 0 & t < 0 \\ \int_0^t 0.01e^{-\frac{x}{100}} & t > 0 \end{cases} = \begin{cases} 0 & t < 0 \\ 1 - e^{-\frac{t}{100}} & t > 0 \end{cases}$$

P(the life time of a randomly selected component is greater than 50) = ?

$$= P(X > 50) = 1 - P(X \leq 50) = 1 - F(50) = e^{-\frac{1}{2}} = \int_{50}^{\infty} f(x) dx$$

# Gamma function

The **gamma function** is denoted  $\Gamma$  and is defined by

$$\Gamma(t) = \int_0^{\infty} x^{t-1} e^{-x} dx \quad t > 0$$

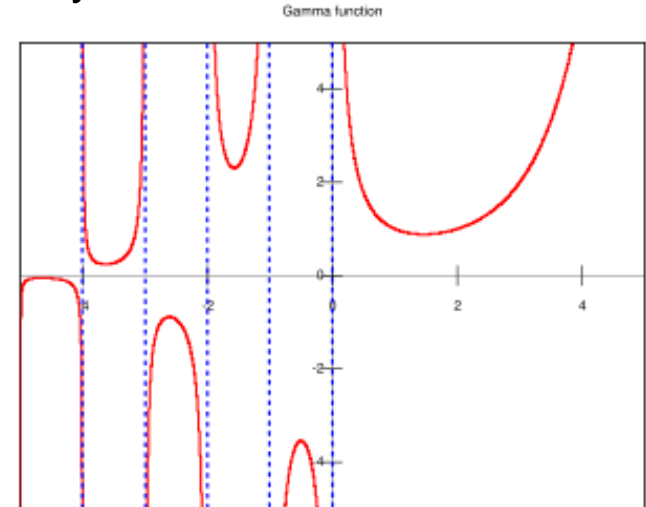
$$\Gamma(1) = \int_0^{\infty} e^{-x} dx = 1$$

It can be proved that  $\Gamma(t) = (t-1)\Gamma(t-1)$

If  $n$  is an integer number then

$$\begin{aligned} \Gamma(n) &= (n-1)\Gamma(n-1) = (n-1)(n-2)\Gamma(n-2) = \\ &= (n-1)(n-2)\dots\dots\dots 1\Gamma(1) = (n-1)! \end{aligned}$$

$$\Gamma(1) = 0! = 1$$





# Gamma distribution

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A continuous random variable  $X$  has **gamma distribution** with parameters  $\alpha$  and  $\beta$  if

$$f(x) = \begin{cases} \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-\frac{x}{\beta}} & x > 0 \\ 0 & x < 0 \end{cases}$$

Find  $E(X)$ ,  $\text{Var}(X)$ ,  $M(t)$  and show that if  $\alpha=1$  then the gamma distribution is an exponential distribution with parameter  $\beta$

# Approximation of Binomial distribution

Reminder:  $X \sim B(n, p)$  ;  $P(X = x) = \binom{n}{x} p^x (1-p)^{n-x}$   
 $E(X) = np$      $V(X) = npq$

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$$X \sim P(\lambda); \quad P(X = x) = e^{-\lambda} \lambda^x / x!$$

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When  $n$  is large then we need to **approximate** the binomial probability

**Theorem** : If  $\underline{n}$  is large and  $\underline{p}$  is small so that  $\underline{np}$  be a constant value, then

$$X \approx P(np) ; \quad P(X = x) \approx \frac{e^{-np} (np)^x}{x!}$$

Discrete approximation